

**Multi-year panel data: Impact of enterprise zones on employment**

**Source:** data file EZUNEM (Wooldridge). 22 cities in Indiana, from 1980 to 1988. Six enterprise zones created in 1984, and 4 more in 1985.

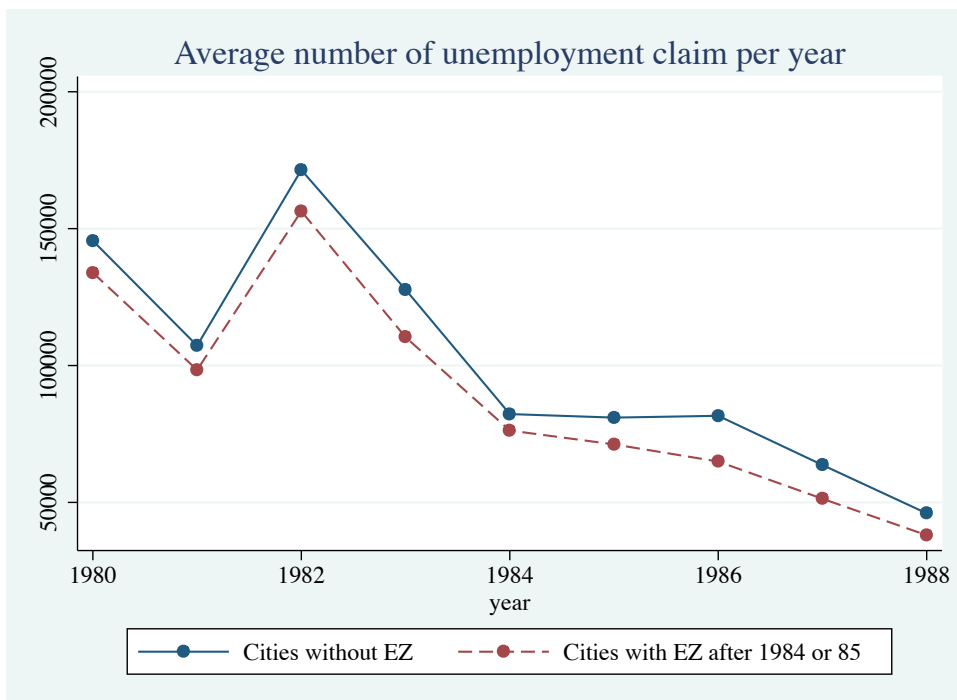
**uclms** is number of unemployment claims file during the year

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. sum
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Variable	Obs	Mean	Std. Dev.	Min	Max
year	198	1984	2.588534	1980	1988
uclms	198	95383.39	89173.63	12360	667208
ez	198	.2323232	.4233845	0	1

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. table year, c(n uclms mean uclms min uclms max uclms)
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year	N(uclms)	mean(uclms)	min(uclms)	max(uclms)
1980	22	140267.2	48490	533598
1981	22	103229	34489	435754
1982	22	164579.3	46225	667208
1983	22	119835.9	34154	545625
1984	22	79541.59	30062	340092
1985	22	76527.23	25004	325797
1986	22	74040.95	21100	303724
1987	22	58051.14	16913	262512
1988	22	42378.32	12360	209103



. reg luclms ez i.year i.city

Source	SS	df	MS	Number of obs =	198
Model	93.7818454	30	3.12606151	F( 30, 167) =	77.75
Residual	6.71437093	167	.040205814	Prob > F =	0.0000
Total	100.496216	197	.510133078	R-squared =	0.9332
				Adj R-squared =	0.9212
				Root MSE =	.20051

luclms	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ez	-.1044144	.0554191	-1.88	0.061	-.2138267	.0049979
year						
1981	-.3216314	.0604572	-5.32	0.000	-.4409903	-.2022725
1982	.1354959	.0604572	2.24	0.026	.016137	.2548548
1987	-.8889487	.0654953	-13.57	0.000	-1.018254	-.7596432
1988	-1.227633	.0654953	-18.74	0.000	-1.356938	-1.098327
city						
2	-.1934845	.099411	-1.95	0.053	-.3897488	.0027797
3	-.3789357	.099411	-3.81	0.000	-.5752	-.1826714
21	.4577925	.0945231	4.84	0.000	.2711782	.6444068
22	.2186422	.099411	2.20	0.029	.0223779	.4149065
_cons	11.67615	.0800792	145.81	0.000	11.51805	11.83425

. xtreg luclms ez i.year, fe i(city)

Fixed-effects (within) regression	Number of obs =	198
Group variable: city	Number of groups =	22
R-sq: within = 0.8416	Obs per group: min =	9
between = 0.0002	avg =	9.0
overall = 0.3528	max =	9
corr(u_i, Xb) = -0.0039	F(9,167) =	98.59
	Prob > F =	0.0000

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1983	-.2192554	.0604572	-3.63	0.000	-.3386144	-.0998965
1984	-.579152	.0623179	-9.29	0.000	-.7021843	-.4561196
1985	-.5917876	.0654953	-9.04	0.000	-.7210931	-.4624821
1986	-.621265	.0654953	-9.49	0.000	-.7505704	-.4919595
1987	-.8889487	.0654953	-13.57	0.000	-1.018254	-.7596432
1988	-1.227633	.0654953	-18.74	0.000	-1.356938	-1.098327
_cons	11.69439	.0427497	273.55	0.000	11.60999	11.77879
sigma_u	.55551507					
sigma_e	.20051387					
rho	.88473197	(fraction of variance due to u_i)				

F test that all u\_i=0: F(21, 167) = 68.95 Prob > F = 0.0000

## Examples of policy analysis with panel data

Carpenter, Christopher, Sabina Postolek, and Casey Warman. 2011. "Public-Place Smoking Laws and Exposure to Environmental Tobacco Smoke (ETS)." *American Economic Journal: Economic Policy*, 3(3): 35-61.

TABLE 2—PUBLIC-PLACE SMOKING LAWS AND EXPOSURE TO ETS IN SPECIFIC PUBLIC PLACES, CTUMS 2005–2008

Exposed to ETS:	On an outdoor patio of a restaurant or bar (1)	Inside a restaurant (2)	Inside a bar or tavern (3)	At a bus stop or shelter (4)	At your workplace (5)	At your school (6)	At any other public place ... † (7)	Anywhere else (8)
Mean of dep. var. in 2005, no law:	0.394	0.539	0.376	0.138	0.274	0.069	0.338	0.044
Public-place smoking law	-0.113** [0.016] {-28.7%}	-0.396** [0.049] {-73.5%}	-0.239** [0.015] {-63.6%}	-0.024 [0.019] {-17.4%}	-0.012 [0.015] {-4.4%}	0.001 [0.005] {1.4%}	-0.051** [0.012] {-15.1%}	-0.030** [0.008] {-68.2%}
Observations	37,693	37,693	37,693	37,693	37,693	37,693	37,693	37,693
R <sup>2</sup>	0.08	0.23	0.15	0.08	0.05	0.19	0.04	0.02
Individual and provincial controls?	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
City and year fixed effects?	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

Notes: Sample includes adults age 18–64 in the CTUMS 2002–2008. The dependent variable in each column is an indicator equal to one if the individual reports being exposed to someone else's smoke in each of the specific areas. See notes to Table 1 for control variables. Robust standard errors in brackets are clustered on statistical areas (SAs). Implied percent impact as a proportion of base year average value of dependent variable for people living in places without a law in curly braces.

† The remainder of the question reads "... such as a shopping mall, arena, bingo hall, concert, or sporting event."

\*\* Significant at the 1 percent level.

\* Significant at the 5 percent level.

Public-place smoking restrictions are the most important non-price tobacco control measures worldwide, yet surprisingly little is known about their effects on exposure to environmental tobacco smoke (ETS). We study these laws in Canada using data with questions about respondents' ETS exposure in public and private places. In fixed effects models we find these laws had no effects on smoking but induced large and statistically significant reductions in public-place ETS exposure, especially in bars and restaurants. We do not find significant evidence of ETS displacement to private homes. Our results indicate wide latitude for health improvements from banning smoking in public places.

Gruber, Jonathan, and Samuel A. Kleiner. 2012. "Do Strikes Kill? Evidence from New York State." *American Economic Journal: Economic Policy*, 4(1): 127-57.

Hospitals now represent one of the largest union sectors of the US economy, and there is particular concern about the impact of strikes on patient welfare. We analyze the effects of nurses' strikes in hospitals on patient outcomes in New York State. Controlling for hospital specific heterogeneity, the results show that nurses' strikes increase in-hospital mortality by 18.3 percent and 30-day readmission by 5.7 percent for patients admitted during a strike, with little change in patient demographics, disease severity or treatment intensity. The results suggest that hospitals functioning during nurses' strikes do so at a lower quality of patient care.

Kapoor, Sacha, and Arvind Magesan. 2014. "Paging Inspector Sands: The Costs of Public Information." *American Economic Journal: Economic Policy*, 6(1): 92-113.

We exploit the introduction of pedestrian countdown signals—timers that indicate when traffic lights will change—to evaluate a policy that improves the information of all market participants. We find that although countdown signals reduce the number of pedestrians struck by automobiles, they increase the number of collisions between automobiles. They also cause more collisions overall, implying that welfare gains can be attained by hiding the information from drivers. Whereas most empirical studies on the role of information in markets suggest that asymmetric information reduces welfare, we conclude that asymmetric information can, in fact, improve it.

TABLE 2—COLLISIONS AND PEDESTRIAN COUNTDOWN SIGNALS

	(1)	(2)	(3)	(4)
Pedestrian countdown Signal activated	-0.055*** (0.006)	-0.022*** (0.004)	0.012** (0.006)	0.011* (0.006)
Controls				
Intersection	No	Yes	Yes	Yes
Month-year	No	No	Yes	Yes
Lagged collisions	No	No	No	Yes
$R^2$	0.001	0.003	0.004	0.003
Intersections	1,794	1,794	1,794	1,794
Observations	107,640	107,640	107,640	105,846

Notes: The dependent variable is number of collisions. Robust standard errors clustered at the intersection level.

\*\*\*Significant at the 1 percent level.

\*\*Significant at the 5 percent level.

\*Significant at the 10 percent level.

TABLE 3—COLLISION INVOLVEMENTS AND CONDITIONS

	Collisions involving:		
	Driver-driver	Driver-pedestrian	Driver-cyclist
Pedestrian countdown Signal activated	0.0117** (0.0052)	-0.0032** (0.0015)	0.0014 (0.0011)
$R^2$	0.004	0.002	0.003
Intersections	1,794	1,794	1,794
Observations	107,640	107,640	107,640

Notes: Robust standard errors clustered at the intersection level. All regressions include fixed effects for the intersection and month-year.

\*\*\*Significant at the 1 percent level.

\*\*Significant at the 5 percent level.

\*Significant at the 10 percent level.